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# The Effects of Active Labour Market Policies for Immigrants Receiving Social Assistance in Denmark

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#### Abstract

We estimate the effect of act ive labour market p rogrammes on the exit rate to r egular employment for non-western immigrants in Denmark who receive social assistance. We use the timing-of-events duration model and rich administrative data. We find large positive postprogramme effects, a nd, surprisingly, even m ost in-programme effects are p ositive. The effects are largest for s ubsidized e mployment pr ogrammes, but effects are a lso large an d significant for direct employment programmes and other programmes. The effects are larger if programmes be gin a fter s ix m onths of unemployment. Implications of o ur e stimates are illustrated b y calculating effects on t he dur ation t o r egular e mployment over a five-year period.

JEL Classification: J64, J24, J68, J61, C41

Keywords: Programme evaluation, duration analysis, timing-of-events model

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#### **1. Introduction**

In m ost E uropean countries e mployment r ates of non -western immigrants are v ery lo w compared to employment rates of natives (OECD, 2005), and this has become a major policy issue because of the important consequences for aggregate labour supply, economic growth and the long-run fiscal sustainability issues. Furthermore, it is often argued that labour market integration of immigrants may be important for social integration and cohesion, and there may be long-term effects through integration of children of immigrants.

In this paper, we investigate the effect of active l abour market programmes (ALMPs) on the duration until regular employment for non-western immigrants in Denmark receiving social assistance (cash b enefits). The reason why we focus on s ocial assistance recipients is that the majority of non-western immigrants in Denmark are not members of an unemployment insurance fund, implying that they are not entitled to unemployment insurance benefits when they become unemployed; instead they receive social assistance, and they are heavily ov er-represented am ong s ocial assistance recipients, e specially a mong l ong-term recipients.

Like ot her recipients of so cial as sistance in D enmark, im migrants a re offered active labour market programmes administered by the job centres. Because a large fraction of non-western immigrants are social assistance recipients, it is very important to assess whether these p rogrammes h ave p ositive effects o n i mmigrants' em ployment p rospects. I f some programmes show positive effects, further targeting of such programmes to immigrants may have large positive effects on l abour market integration and, indirectly, on social integration in general.

We use r ich ad ministrative data covering all immigrants in Denmark, and we employ the timing-of-events duration model of Abbring and van den Berg (2003).<sup>1</sup> We find large pos itive pos t-programme ef fects, an d, s urprisingly, most i n-programme ef fects are positive too. We find the largest effects for subsidized employment programmes, but effects are also l arge an d significant f or d irect em ployment p rogrammes as w ell as for other programmes. Effects are larger for programmes beginning after six months of unemployment. Implications of our estimates are illustrated by calculating marginal effects on the expected duration t o r egular e mployment over a f ive-year pe riod f or t ypical starting t imes an d programme durations. Subsidized employment programmes shorten the social assistance spell by a bout 10 m onths for women and 15 m onths for men. The effect of direct employment programmes is 3.7 months for women and 2.6 months for men.

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Only very few papers have investigated the effects of ALMPs specifically for immigrants. Clausen *et al.* (2009) study effects of programmes offered to *newly arrived* non-western immigrants in Denmark. They find that post-programme effects on the hazard rate to regular employment are significantly positive for wage subsidy programmes, but not for other types of programmes. Bolvig *et al.* (2003) reach a similar conclusion investigating effects of ALMPs for so cial assistance recipients in the second largest municipality in Denmark, but they do not estimate separate effects for natives and immigrants.

The overall finding in previous studies on programmes for unemployed workers is that earnings effects as well as employment effects are small; see the surveys in Stanley et al. (1999), Heckman et al. (1999), Kluve (2006), and Card et al. (2010). Stanley et al. (1999) summarise the effects of several US programmes. They find that temporary wage subsidy programmes increase the probability of finding jobs in the subsidy period, but not in the long run, and that job search courses and counselling shorten unemployment duration. The surveys in bot h S tanley et al. (1999) a nd H eckman et al. (1999) c onclude t hat e ffects a re heterogeneous: some ALMPs may have positive effects for unemployed with a weak position in the labour market, but for other groups the effects are smaller if at all positive. Our findings of rather strong positive effects for non-western immigrants in Denmark are consistent with these previous results, since these immigrants in general have a weak position in the Danish labour market because of language problems and a low level of education or non-transferable education from their country of origin. The findings in the surveys of European labour market programmes by Kluve (2006) and Card et al. (2010) are similar to those for the US: 'privatesector i ncentive pr ogrammes' (including w age s ubsidies) a nd 'services a nd sanctions' (including j ob s earch assistance, c ounselling, and s anctions f or non -compliance) ty pically have positive em ployment ef fects, whereas 'traditional labour market training' (including classroom training and on-the-job training) has much smaller and often insignificant effects, and 'direct employment programmes in the public sector' are rarely effective and often have negative employment effects. Card et al. (2010) do show, however, that these negative effects tend to become smaller when effects are studies over a longer time perspective.

This paper is organized as follows. Section 2 describes the institutional setting. Section 3 develops the econometric model. Section 4 describes the data used in the empirical analysis. Section 5 presents the results. Section 6 concludes and discusses policy implications.

#### 2. Institutional setting

In Denmark, membership of an unemployment insurance fund is voluntary, and membership of a given fund may be restricted by education and previous employment. Only members will receive U I be nefits i ft hey be come une mployed. N on-insured workers w ho be come unemployed may i nstead r eceive so cial assistance b enefits w hich ar e means-tested. A s discussed a bove, w e f ocus on no n-insured i mmigrants in t his p aper b ecause there are considerably m ore non -western im migrants w ho a re n on-insured t han insured, a nd, in particular, immigrants are over-represented among long-term recipients of so cial assistance. Therefore, t he effect of pr ogrammes a imed a t l abour m arket i ntegration of non -western immigrants receiving social assistance is of particular interest.

Social assistance recipients will receive ALMP offers from the local authorities, and t hey ha ve t o participate in s uch a pr ogramme in order to remain e ligible f or s ocial security benefits, which are of unlimited duration. A wide range of ALMPs are being used. In this paper we distinguish between three types: employment with a wage subsidy (mainly in the private sector), direct employment pr ogramme (mainly in the public sector), and other programmes which include education, training, and counselling programmes.

According t o na tional 1 aw, pe rsons s hould be of fered pa rticipation i n a programme within 12 months from the beginning of the social-assistance spell.

#### 3. Econometric model

We use the timing-of-events duration model of Abbring and van den Berg (2003). We model the duration of the social assistance spell to regular employment si multaneously with the duration from the beginning of the social assistance spell until entry into active labour market programmes. The duration of the social assistance spell includes the time spent in ALMPs. The model takes a ccount of non-random selection i nto these programmes with r espect to observable and unobservable covariates. A ssuming mixed proportional hazard (MPH) rates and no an ticipation effects, the treatment effects (i.e. the effects of participating in labour market programmes) are non-parametrically identified; see Abbring and van den Berg (2003). The no-anticipation assumption requires that a treatment starting at time t should not affect the outcome state (employment or non-employment) before time t. This may be a reasonable assumption in the present application since typically social assistance recipients are not able to p redict n either t ime of tre atment n or the s pecific p rogramme to w hich th ey m ay b e assigned. First of all, municipalities have a large degree of discretion regarding the timing of treatment, so me o ffer programmes at a very early stage of so cial assistance receipt, while others typically do not meet the 12-month criterion. In addition, there is a significant variation regarding a ssignment to A LMPs b etween case w orkers w ithin m unicipalities, and f inally programme participation d ecisions are taken at meetings with case w orkers, the timing of which is also random to some extent. Thus, the starting times vary a lot for each category of programmes, and f or a ll t ypes of pr ogrammes t aken t ogether t he s tandard de viation of duration until treatment is 15 months; see Table 1 (in the next section).

Given the no-anticipation and MPH assumptions, no e xclusion restrictions are needed. The intuition behind the identification strategy intrinsic to the timing-of-events model is that there is some r andom variation in the duration until treatment. This enables a separation between the treatment effect and the unobserved heterogeneity, which is assumed time-invariant.

The baseline hazard rate is modelled flexibly by a piecewise constant function, and time-varying variables are used for modelling in- and post-programme effects of ALMPs.

Normalising the time at which immigrants begin a social assistance spell to zero, the non -negative st ochastic v ariables  $T_u$  and  $T_p$  measure duration u ntil e mployment a nd duration until programme participation, respectively. By construction,  $T_p \pounds T_u$ . If  $T_p < T_u$ , the immigrant participates i n a pr ogramme, a nd i f  $T_p = T_u$ , he does not (the duration until participation is right-censored).

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Let x(t) be a v ector of o bserved t ime-varying variables, a ndl et  $n_u$  and  $n_p = (n_{p1}, ..., n_{pJ})$  denote unobserved variables affecting the hazard rates to employment and to participation in e ach of the J programmes, respectively. The hazard rate to participation in programme j is

$$\boldsymbol{q}_{pj}\left(\boldsymbol{t}_{p} \mid \boldsymbol{x}\left(\boldsymbol{t}_{p}\right), \boldsymbol{n}_{pj}\right) = \boldsymbol{I}_{pj}\left(\boldsymbol{t}_{p}\right) \exp\left(\boldsymbol{x}\left(\boldsymbol{t}_{p}\right)\boldsymbol{b}_{pj} + \boldsymbol{n}_{pj}\right)$$
(1)

where  $I_{pj}(t_p)$  are piecewise constant baseline hazard rates,

$$I_{pj}(t_{p}) = \mathop{\text{a}}\limits_{m=1}^{M} g_{pjm} \mathbb{1}\{h_{m-1} \notin t_{p} < h_{m}\}, \quad h_{0} = 0, h_{M} = \forall, \ h_{0} < \mathsf{K} < h_{M}.$$
(2)

In t his a pplication, where t he t ime unit is months,  $h_1 = 3$ ,  $h_2 = 6$ ,  $h_3 = 12$ ,  $h_4 = 24$ . The hazard r ate t o p rogramme p articipation i s t he su m o f t he h azard r ates to t he sp ecific programmes:

$$\boldsymbol{q}_{p}\left(\boldsymbol{t}_{p} \mid \boldsymbol{x}\left(\boldsymbol{t}_{p}\right), \boldsymbol{n}_{p}\right) = \overset{J}{\underset{j=1}{\overset{J}{\mathbf{a}}}} \boldsymbol{q}_{pj}\left(\boldsymbol{t}_{p} \mid \boldsymbol{x}\left(\boldsymbol{t}_{p}\right), \boldsymbol{n}_{pj}\right).$$
(3)

Participation in the *J* different programmes is denoted by two time-varying 1'*J*-dimensional vectors of dummy variables,  $d_1(t)$  and  $d_2(t)$ . The *j*th element of  $d_1(t)$  is equal to 1 if the individual is participating in programme *j* at time *t*, and 0 otherwise; at most, one element of  $d_1(t)$  is equal to 1 at time *t*. Similarly, the *j*th element of  $d_2(t)$  is equal to 1 if the individual participated in programme *j* before time *t*, but is no longer participating, and 0 otherwise. We only c onsider effects of t he first p rogramme d uring a so cial assistance sp ell; if a p erson participates i n a s econd pr ogramme, t he obs ervation i s r ight-censored at t he time t his participation begins.

The exit rate to employment is given by

$$q_{u}(t_{u} | x(t_{u}), d_{1}(t_{u}), d_{2}(t_{u}), n_{u}) = I_{u}(t_{u}) \exp(x(t_{u}) b_{u} + d_{1}(t_{u}) d_{1} + d_{2}(t_{u}) d_{2} + n_{u})$$
(4)

where the baseline hazard  $l_u(t_u)$  has a form similar to (2), and b,  $d_1$  and  $d_2$  are vectors of parameters; b is the effect of the control variables,  $d_1$  the in-programme effect, and  $d_2$  the post-programme effect after completed programme participation. The model takes account of endogeneity of  $d_1(t)$  and  $d_2(t)$  through possible correlation between the unobserved variables  $n_u$  and  $n_p$ .

Let c be a dummy variable equal to 1 if the person exits to employment and 0 otherwise; similarly, let  $c_j = 1$  if the person participates in programme j. Then the contribution to the likelihood function of a specific spell, given observed variables, is

The likelihood function is

$$L = \overset{\bullet}{\mathbf{\omega}} L(n_u, n_p) dF(n_u, n_p)$$
(6)

where F is the distribution function of  $(n_u, n_p)$ . We take account of the fact that a g iven person may have more than one social assistance spell. To simplify the estimation, we apply a discrete distribution (Heckman and Singer, 1984). Specifically, we assume that  $(n_u, n_p)$  has a discrete distribution with 2'2 mass points. This implies that the unobserved components of the s election in to the J different p rogrammes ar e p erfectly correlated, b ut the correlation between  $n_u$  and  $n_p$  is unrestricted.

#### 3.1 Marginal effects on the hazard rate and on the duration to employment

The marginal effects of the control variables on the hazard rate to employment are given by the coefficients  $b_u$  (ignoring the effects via programme participation). Thus, the coefficient of the  $h^{\text{th}}$  explanatory variable,  $b_{uh}$ , is equal to the change in the logarithm of the hazard rate to employment when this variable is changed by 1 unit holding all other variables constant. Similarly,  $d_1$  and  $d_2$  are th e m arginal in-programme and p ost-programme effects, respectively, o f p articipation in labour m arket programmes on t he log hazard r ate t o employment.

The t otal e ffect of participation in a specific programme on the expected duration until employment depends of c ourse on  $d_1$  and  $d_2$ , but a lso on t he timing and duration of the programme and on the basic level of the hazard rate to employment (since  $d_1$ and  $d_2$  affect the hazard rate multiplicatively) determined by individual characteristics. We calculate the marginal effects of programme participation for a 'reference person' g iven a range of t ypical p rogramme st arting t imes a nd d urations. T hese marginal effects ar e calculated a s t he d ifference i n expected m ean d uration t o em ployment w ith an d w ithout programme participation. In these calculations we use restricted mean durations, i.e. expected mean durations up to a predetermined endpoint,  $T_{max}$  (which is taken to be five years).

Let  $z(t) = (x(t), d_1(t), d_2(t))$  denote the vector of covariates. Let  $z^1(t)$  denote the covariates when an individual is assigned to a given programme at a given time and with a given duration, and let  $z^0(t)$  denote covariates when the individual is not assigned to any programme. Then the marginal effect of t his programme is d effined as t he d ifference in restricted expected durations:

where the restricted expected duration is the expected area under the survivor curve up to time  $T_{max}$ 

$$E_{T_{\text{max}}}\left[T_{u} \mid z(t)\right] = E \stackrel{\acute{e}}{\underline{e}} \overset{T_{\text{max}}}{Q} S(r) dr \overset{\grave{v}}{\underline{d}}$$
(8)

and the survivor function is calculated from the hazard rate:

$$S(t) = \exp\left(-\grave{\mathbf{Q}}^{t}\boldsymbol{q}_{u}(s \mid z(s))ds\right).$$
(9)

Details on the calculation of marginal effects are described in Section 5.

#### 4. Data

We u se d ata f rom sev eral ad ministrative r egisters, w hich ar e co llected an d merged b y Statistics Denmark. Our data set covers the period 1984-2004 and contains all immigrants in Denmark, but we restrict the analysis to immigrants from non-western countries.<sup>2</sup>

We further restrict the sample to p ersons who were r esidents in D enmark in 1997 or 1998 due to the fact that there was a major institutional change in 1999 a ffecting newly arrived immigrants, and also that the administrative data on ALMP participation for immigrants who g of their residence p ermit in 1 999 a re incomplete. Also, we re strict the analysis to immigrants who be gan a social assistance spell b etween Jan uary 1 997 and December 2003 due to the fact that data on participation in ALMPs are only available from 1997 onw ards. F inally, we re strict the analysis to so cial assistance spells where the immigrants were 18-66 years of age when the spell began.

The final estimation sample contains 66,768 social assistance spells, 31,215 for women and 35,553 for men; see Table 1.

The upper panel of Table 1 shows figures for both women and men. For 38% of the so cial a ssistance sp ells, the immigrant is participating in an activation programme. Of those who participate in programmes, 7.4% concerns employment with a wage subsidy, and 26.7% direct employment programmes. Durations of each type of activation programme vary a lot (mean duration is 5-6 months with a standard deviation of 4-6 months), but one reason for this is that some programmes are interrupted, for instance because the participant finds a job. T here is a lso a large variation in t he t ime of e ntry i nto pr ogrammes: O n a verage, immigrants who participate in a programme enter about 13 months after the start of the social assistance spell, and the standard deviation is 15 m onths. 34% of all the immigrant s ocial assistance spells have employment as d estination state. The percentage is 39 for spells not involving ALMP participation and 26 for spells with ALMP participation, varying from 51% for employment with wage subsidy to 20 for 'other programmes'.

#### [See Table 1]

The two lower panels in Table 1 show figures for women and men, separately. The fraction participating in employment programmes is larger for men, whereas the fraction participating in 'o ther p rogrammes' is la rger f or women. The a verage duration of di rect e mployment programmes and 'other programmes' is about 1 m onth longer for women than for men, and women are enrolled in these programmes much later than men (about 3 and 6 months later,

respectively). The probability of employment as the destination state is considerably larger for men than for women, especially for social assistance spells not involving ALMP participation and for spells with participation in 'other programmes'.

Durations of t he so cial ass istance sp ells ar e i n g eneral v ery long and considerably longer for women than for men. This is illustrated in Figure 1 by non-parametric Kaplan-Meier survival functions. For women, the probability that a social as sistance sp ell lasts at least 12, 24, 48 and 84 m onths is 79, 71, 61 and 54%, r espectively. For men the corresponding percentages are 67, 54, 43 and 37.

#### [See Figure 1]

We est imate sep arate models f or men a nd women. Control variables i n t he an alyses a re measured in the year where the social assistance spell began. Controls are variables for years since m igration, c ountry of or igin, type o f r esidence p ermit, age, whether t he pe rson ha s children in di fferent a ge gr oups, whether t he pe rson i s single, w orking experience in Denmark, type of municipality, education, h ealth indicators based on t he number of yearly doctor visits, the local unemployment rate in the commuting area of the municipality, and the calendar year in which the social assistance spell began. Table A1 in the Appendix contains descriptive statistics for the control variables.

#### 5. Results

Table 2 shows the main estimation results. The first columns show the estimation result for women, and the last columns the result for men. Table 2 shows results for parameters related to ALMP participation only. Estimated parameters of other explanatory variables are shown in the Appendix, Table A2. For each of the three types of ALMPs, Table 2 shows parameter estimates f or in-programme and pos t-programme effects. F urthermore, w e al low t hese parameters to differ according to whether the immigrant entered the programme earlier than six months after the start of the social assistance spell, or later. Thus, the table shows for each gender six estimates of in-programme effects are positive and clearly significant. Employment with a wage subsidy has the largest effect followed by direct employment programmes. For all three types of ALMPs, the post-programme effects on the hazard rate to employment are larger if the activation period begins after six months (rather than during the first six months of the social assistance spell).

#### [See Table 2]

Somewhat surprisingly, most in-programme effects are also positive, indicating that ALMP participation increases the hazard rate to ordinary employment also during participation. The standard f inding he re i s one of 1 ock-in e ffects during programme participation. The only exception is for subsidized employment spells which begin less than six months after the start of the so cial ass istance sp ell; here we do f ind a 1 ock-in effect, but it is only marginally significant. Clausen *et al.* (2009) also find positive, but insignificant, in-programme effects for subsidized employment programmes for n ewly a rrived i mmigrants. Another s urprising feature of t he est imates i s t hat t he positive in-programme effect i s 1 arger t han t he postprogramme effect for direct employment programmes and 'other programmes' be ginning at least six months after the start of the social assistance spell (and for males the two types of effects ar e about eq ual i n si ze f or direct employment a nd 'other' programmes be ginning within the first six months of the social assistance spell).

One reason why the estimated in-programme effects are positive and rather large may be that immigrants on social assistance benefits in general have weak qualifications, including weak host country language proficiency, and only a limited know ledge of the Danish labour market. Participating in ALMPs may therefore be particularly important for immigrants' employment chances, also because participation may serve as a positive signal to employers who may have more difficulties assessing the qualifications of immigrants than of natives.

The estimated parameters – and thereby the relative effects on the hazard rate to employment – are generally larger for women than for men, but the initial level of the hazard rate to employment is considerably lower for women than for men. The post-programme parameter of subsidized employment starting after six months for women is 2.4 indicating a tenfold increase in the hazard rate to employment. The corresponding parameter for men (1.9) indicates a fivefold increase in the hazard rate. The corresponding in-programme parameters of 1.3 a nd 0.84 i ndicate i ncreases in the hazard rate to employment of 277% and 132%, respectively.

Table 3 shows the marginal effects of participation in activation programmes on the restricted mean duration to employment over a five-year period, calculated from a large number of typical programme spells. The effects are calculated for a 'reference person'. The characteristics chosen for this person (which a ffect the size of the marginal effects, but not their sign or statistical significance) ar e g iven b y the r efference ca tegories of e ach s et o f categorised variables (less than 5 years since migration, from former Yugoslavia, refugee, age above 45 ye ars, has children be low 2 ye ars of a ge (but no older children), cohabiting, no working experience in Denmark, lives in a big city, education unknown, 1-4 visits to the doctor, t he so cial assistance sp ell b egan i n 1 997) and by t he a verage of t he l ocal unemployment r ate. If s uch a person di d not participate i n a ny activation programme, the restricted mean duration to employment would be 55 months for a woman and 42.4 months for a man (see the first row in Table 3). These numbers are very large; the maximum would be 60 months (given that it is restricted to be at most 5 years).

Given the v alues of the c ontrol v ariables, the r estricted m ean dur ation w hen participating in a given programme depends on the duration until entering the programme and the duration of the programme. Starting time and duration of a given programme vary a lot (see T able 1). T herefore, w e cal culate 'the' m arginal effect of a g iven p rogramme as an average over several typical variants of the programme defined by starting time and duration. Specifically, we use the following simplifying assumptions. Programme spells may have three different s tarting times and durations defined by the first, s econd and t hird quartile in the observed di stributions of s tarting times and durations (for women and men, respectively). These quartiles are shown in Table A3 in the Appendix. Thus, there are nine different types of a given programme, and it is assumed that each type has equal probability 1/9.

Participation in ALMPs leads to a large reduction in the restricted mean duration to employment as shown in Table 3. Employment with a wage subsidy shortens the restricted mean duration by 1 0 months for women and by 15 months for men, which are very large effects. The effects of direct employment programmes and 'other ALMPs' are smaller, but also large; they r educe restricted m ean durations by a bout 4 and 2 m onths, r espectively. Again, the estimated effects are larger for men than for women.

#### [See Table 3]

The est imated m arginal effects in Table 3 are clearly significant. Standard e rrors of the marginal effects ar e c alculated f rom simulations of t he est imated p arameters. G iven t he estimated parameters and their estimated covariance matrix, we draw 500 random parameter vectors, and calculate the marginal effects for each parameter vector. The estimated standard error of the marginal effect of a g iven pr ogramme is the standard d eviation of the 500 calculated marginal effects of this programme.

To our knowledge, only one earlier study has focused on effects of ALMPs for immigrants receiving social assistance. Thus, Clausen *et al.* (2009) analysed effects for *newly arrived* immigrants in Denmark and found significant effects for subsidized employment (a reduction in the restricted mean duration to employment of a bout four months over a fouryear period), but not for direct employment programmes or other programmes. Comparing our results to the broader literature on effects of ALMPs on labour market outcomes (see the surveys in Stanley *et al.*, 1999; Heckman *et al.*, 1999; Kluve and Schmidt, 2002; Kluve, 2006; Card *et al.*, 2010), the positive effect of employment with a wage subsidy on the hazard rate to regular employment is consistent with most previous studies. Our finding of significant and rather la rge effects f or d irect em ployment p rogrammes is l ess c onsistent with previous findings, since m ost st udies f ind only sm all and o ften i nsignificant effects. Our cat egory 'other programmes', for which we find a small but significant effect, includes very different programmes su ch as t raining, sp ecial em ployment pr ogrammes and c ounselling. Most previous studies find po sitive effects of c ounselling, m arginally p ositive effects of sp ecial employment programmes and small positive or insignificant effects of training.

#### 6. Discussion and conclusion

Using the timing of events duration model and a large administrative data set, this study finds large and substantial positive effects of having participated in ALMPs on the hazard rate to regular e mployment for i mmigrants re ceiving s ocial a ssistance in D enmark. We e stimate separate effects for three categories of programmes, and we allow both in-programme and post-programme effects to differ according to whether the program began before or after 6 months. The post-programme effects are largest for subsidized employment programmes. We also f ind p ositive in-programme effects except f or s ubsidized e mployment programmes starting early. Both lock-in and post-programme effects are much larger when the program begins at least six months after the start of the social assistance spell.

The total effect of ALMP participation depends on both in-programme and postprogramme effects. To assess the overall effect, we calculate the marginal effects on the mean duration to regular employment over a five-year period, given a range of typical starting times and durations of ALMP sub-spells. Subsidized employment programmes reduce the duration of so cial a ssistance sp ells b y 1 0-15 months, d irect e mployment programmes r educe it by about 4 months, and other programmes by about 2 months.

The finding that subsidized employment is the most effective type of ALMP is consistent with an earlier study focussing on newly arrived immigrants and with other studies of ALMP effects for unemployed in general. However, this type of ALMP is by far the least frequently a pplied i n D enmark; onl y 7.4% of ALMPs of fered t o i mmigrants on s ocial assistance are subsidized employment programmes (4.9% for females and 9.6% for males). Our re sults thus indicate that labour market i ntegration of i mmigrants may b e im proved considerably by targeting subsidized employment programmes to unemployed immigrants.

Since w e a lso est imate considerable positive effects of d irect employment programmes and other programmes for immigrants – effects which are larger than similar effects for unemployed in general found in other studies – offering these programmes to a larger number of unemployed immigrants may also be beneficial to labour market integration and labour supply in general.

However, e ven though our r esults indicate l arge b eneficial ef fects o f an intensified use of ALMPs for immigrants, it may not be optimal to offer these programmes at a v ery early st ate of t he u nemployment/social assistance spell s ince o ur est imates i ndicate much larger positive effects on the hazard rate to employment if the ALMPs begin after six months of social assistance r eceipt. Such a de cision r equires w eighing t he c osts of programmes against the saved social assistance.

In conclusion, we have found very large positive employment effects of ALMPS for non-western immigrants, suggesting a much more active role for activation policies in national s trategies for integrating i mmigrants i nto the c ountry. A lso, we would definitely advocate more research on the topic.

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#### **Dansk Sammenfatning**

#### Effekten af aktivering for ikke-vestlige indvandrere på kontanthjælp

Formålet m ed d ette p apir er at e stimere e ffekten a f aktivering p å v arigheden t il o rdinær beskæftigelse for ikke-vestlige indvandrere på kontanthjælp. Der tages højde for, at det ikke er ti lfældigt, h vem d er b liver a ktiveret i b estemte foranstaltninger (v ed b rug a f tim ing-ofevents-metoden). Analysen er baseret på registerdata for kontanthjælpsforløb i perioden 1997-2004 for indvandrere, der kom til Danmark før 1999. R esultaterne viser, at aktivering har en stor pos itiv e ffekt på i ndvandrernes s andsynlighed f or a t komme i or dinær be skæftigelse. Effekterne er størst for ansættelse med løntilskud, men de er også store for virksomhedspraktik og a nden aktivering. Effekterne er noget større for mænd end for kvinder. Den gennemsnitlige forventede varighed på kontanthjælp inden for en femårsperiode er ca. 42 måneder for mænd. Ansættelse med løntilskud afkorter denne periode med ca. 15 måneder, virksomhedspraktik afkorter den med ca. 4½ måned og a nden aktivering afkorter den ca. 2½ måned. For kvinder er de tilsvarende reduktioner i varigheden på kontanthjælp henholdsvis ca. 10, 3½ og  $1\frac{1}{2}$  måned.

### Appendix

Tables A1 and A2 show descriptive statistics and parameter estimates, respectively, for control variables. Table A3 shows characteristics of ALMP spells used to calculate marginal effects, namely the quartiles in the distributions of duration of social assistance spells until start of programme and time spent in the programme, respectively.

[See Table A1] [See Table A2] [See Table A3]

	Social assistance		ALMP Duration of ALMP		Duration of spell until		Spells with		
	sp	pells	spells	subspells (months)		start of ALMP (months)		employment destination	
	Ν	per cent	per cent	Mean	SD	Mean	SD	Ν	per cent
Females and males									
Social assistance spells with ALMPs	25,541	38.3	100.0	5.8	6.1	13.4	15.3	6,664	26
- Employment with wage subsidy	1,893	2.8	7.4	5.3	4.2	8.6	10.2	959	51
- Direct employment programme	6,814	10.2	26.7	5.7	6.0	11.2	13.1	2,273	33
- Other programmes	16,834	25.2	65.9	5.9	6.3	14.8	16.4	3,432	20
Social assistance spells with no ALMP	41,227	61.7						15,916	39
All	66,768	100.0		5.8	6.1	13.4	15.3	22,580	34
Females									
Social assistance spells with ALMPs	11,802	37.8	100.0	6.3	6.3	16.2	17.5	2,575	22
- Employment with wage subsidy	573	1.8	4.9	5.4	4.2	8.5	10.1	278	49
- Direct employment programme	2,958	9.5	25.1	6.2	6.1	13.1	14.9	953	32
- Other programmes	8,271	26.5	70.1	6.4	6.5	17.9	18.4	1,344	16
Social assistance spells with no ALMP	19,413	62.2						5,984	31
All	31,215	100.0		6.3	6.3	16.2	17.5	8,559	27
Males									
Social assistance spells with ALMPs	13,739	38.6	100.0	5.3	5.9	10.9	12.7	4,089	30
- Employment with wage subsidy	1,320	3.7	9.6	5.2	4.3	8.7	10.2	681	52
- Direct employment programme	3,856	10.8	28.1	5.3	5.9	9.8	11.3	1,320	34
- Other programmes	8,563	24.1	62.3	5.3	6.0	11.7	13.6	2,088	24
Social assistance spells with no ALMP	21,814	61.4						9,932	46
All	35,553	100.0		5.3	5.9	10.9	12.7	14,021	39

Table 1. Descriptive statistics of immigrants' social assistance spells and ALMP participation

	Females			Males		
	Coeff.	SE		Coeff.	SE	
ALMP spell begins less than 6 months	s after start	of social	assistan	ce spell		
Lock-in effects:						
Employment with wage subsidy	-0.2394	0.1505		-0.1887	0.0917	**
Direct employment programme	0.4131	0.0702	***	0.3791	0.0537	***
Other ALMPs	0.1693	0.0616	***	0.1588	0.0446	***
Post-programme effects:						
Employment with wage subsidy	1.4287	0.1333	***	1.2430	0.0805	***
Direct employment programme	0.6844	0.0920	***	0.3757	0.0718	***
Other ALMPs	0.2369	0.0769	***	0.1240	0.0570	**
ALMP spell begins at least 6 months after start of social assistance spell						
Lock-in effects:						
Employment with wage subsidy	1.3192	0.1786	***	0.8382	0.1100	***
Direct employment programme	1.5442	0.0695	***	1.0925	0.0607	***
Other ALMPs	0.9390	0.0563	***	0.7245	0.0505	***
Post-programme effects:						
Employment with wage subsidy	2.4127	0.1235	***	1.8976	0.0738	***
Direct employment programme	1.2847	0.0871	***	0.4565	0.0861	***

Table 2. Estimates of effects of participation in ALMPs on the hazard rate to employment

Note: \*\*\*, \*\* and \* indicate significance at the 1%, 5% and 10% level, respectively.

0.5277 0.0701

Other ALMPs

Parameter estimates for the other explanatory variables and for the duration dependent constant terms are shown in the Appendix, Table A2.

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0.3266 0.0607

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Table 3. Marginal effects of labour market programmes: Change in restricted mean duration of social assistance spells over a five-year period (measured in months)

	F	emales				
	Restricted	Marginal		Restricted	Marginal	
	mean	effect	SE	mean	effect	SE
	duration			duration		
No ALMP	55.0			42.4		
Employment with wage subsidy	45.1	-9.9	1.3	27.3	-15.1	1.0
Direct employment programme	51.4	-3.7	0.5	37.8	-4.6	0.6
Other ALMPs	53.5	-1.5	0.2	39.8	-2.6	0.4

Note: The calculation of marginal effects is described in the text.

	Females				Males				
Variable	Mean	SD	Min	Max	Mean	SD	Min	Max	
Years since migration: 5-9	0.248	0.432	0	1	0.232	0.422	0	1	
Years since migration: 10-19	0.183	0.386	0	1	0.235	0.424	0	1	
Years since migration: 20-	0.166	0.372	0	1	0.176	0.381	0	1	
From Turkey	0.151	0.358	0	1	0.114	0.318	0	1	
From Europe (except former Yugoslavia)	0.090	0.286	0	1	0.082	0.274	0	1	
From Africa (except Somalia)	0.071	0.256	0	1	0.081	0.273	0	1	
From Somalia	0.096	0.294	0	1	0.120	0.325	0	1	
From American countries	0.021	0.144	0	1	0.017	0.129	0	1	
From Afghanistan	0.018	0.135	0	1	0.023	0.148	0	1	
From Iraq	0.076	0.264	0	1	0.118	0.322	0	1	
From Iran	0.050	0.218	0	1	0.063	0.243	0	1	
From Sri Lanka	0.037	0.188	0	1	0.025	0.155	0	1	
From other Asian countries	0.088	0.284	0	1	0.044	0.205	0	1	
From Pakistan	0.043	0.203	0	1	0.037	0.188	0	1	
From Lebanon	0.059	0.236	0	1	0.087	0.282	0	1	
From other countries	0.041	0.199	0	1	0.043	0.203	0	1	
Family reunified to a refugee	0.128	0.334	0	1	0.044	0.204	0	1	
Family reunified to non-refugee	0.169	0.375	0	1	0.099	0.298	0	1	
EU residence permit	0.035	0.183	0	1	0.020	0.140	0	1	
Unknown type of residence permit	0.313	0.464	0	1	0.343	0.475	0	1	
Age 16-24	0.215	0.411	Ō	1	0.189	0.391	Ō	1	
Age 25-34	0.400	0.490	0	1	0.378	0.485	0	1	
Age 35-44	0.264	0.441	0	1	0.309	0.462	0	1	
Children 3-6 years of age	0.194	0.395	Ő	1	0.122	0.327	Ő	1	
Children 7-17 years of age	0 205	0 404	Ő	1	0 1 2 3	0 329	Ő	1	
No children	0.321	0 467	Ő	1	0.531	0 4 9 9	Ő	1	
Single	0.331	0.471	Ő	1	0.423	0.494	Ő	1	
Working experience up to 1 year	0.224	0.417	Ō	1	0.279	0.448	Ō	1	
Working experience 1-3 years	0 1 1 2	0.315	Ő	1	0 175	0 380	Ő	1	
Working experience 3 years or more	0.064	0 245	Ő	1	0 1 2 0	0 325	Ő	1	
Lives in provincial town municipality	0 343	0.475	Ő	1	0 326	0.469	Ő	1	
Lives in a rural district	0 1 1 3	0 317	Ő	1	0.097	0 296	Ő	1	
Danish education	0 166	0.372	Ő	1	0.217	0.413	Ő	1	
Years of Danish education	1 694	3 858	0 0	20	2 205	4 281	0 0	20	
Foreign education	0.346	0 476	0	-0	0.325	0.468	0	-0	
Vears of foreign education	3 876	5 656	0 0	18	3 957	5 962	0 0	18	
No visits to doctors	0 1 50	0 357	0	10	0 241	0.428	0	1	
5-9 visits to doctors	0.270	0.327	0 0	1	0.207	0.405	0 0	1	
10-19 visits to doctors	0.270	0.414	0	1	0.108	0.100	0	1	
20 or more visits to doctors	0.076	0.265	0	1	0.025	0.156	0	1	
Local unemployment rate x 10	0.670	0.150	0 326	1 361	0.622	0.150	0 317	1 361	
Social assistance spell began 1998	0.192	0.150	0.520	1.501	0.194	0.190	0.517	1.501	
Social assistance spell began 1999	0.172	0.394	0	1	0.194	0.393	0	1	
Social assistance spell began 2000	0 1 1 0	0.302	0	1	0 1 1 1	0.373	0	1	
Social assistance spell began 2000	0 108	0 3 1 1	0	1	0 105	0.317	0	1	
Social assistance spell began 2001	0.108	0.207	0	1	0.105	0.307	0	1	
Social assistance spell began 2002	0.090	0.287	0	1	0.080	0.271	0	1	

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Note. R eference categories are: Y ears since migration less than five y ears; from former Y ugoslavia; refugee; 45-66 years of age; children 0-2 years of age; married or cohabiting; no working experience (in Denmark); lives in the metropolitan area of Copenhagen or one of the three largest provincial cities in Denmark; education unknown; 1-4 visits to doctors; spell began in 1997

	Females				Males		
	Coeff.	SE		Coeff.	SE		
Years since migration: 5-9	0.276	0.038	***	0.134	0.029	***	
Years since migration: 10-19	0.008	0.053		-0.102	0.040	***	
Years since migration: 20-	-0.126	0.050	***	-0.092	0.038	***	
From Turkey	-0.455	0.063	***	-0.228	0.049	***	
From Europe (except former Yugoslavia)	-0.353	0.060	***	-0.322	0.048	***	
From Africa (except Somalia)	-0.360	0.067	***	-0.267	0.051	***	
From Somalia	-0.971	0.071	***	-0.532	0.048	***	
From American countries	-0.176	0.094	*	-0.275	0.092	***	
From Afghanistan	-0.574	0.124	***	-0.249	0.076	***	
From Iraq	-0.851	0.076	***	-0.458	0.044	***	
From Iran	-0.253	0.075	***	-0.176	0.052	***	
From Sri Lanka	0.226	0.077	***	0.057	0.065		
From other Asian countries	-0.119	0.062	*	-0.170	0.056	***	
From Pakistan	-0.676	0.084	***	-0.135	0.063	**	
From Lebanon	-0.998	0.084	***	-0.485	0.052	***	
From other countries	-0.588	0.082	***	-0.388	0.056	***	
Family reunified to a refugee	-0.262	0.058	***	-0.110	0.063	*	
Family reunified to non-refugee	0.074	0.052		0.117	0.044	***	
EU residence permit	0.082	0.079		-0.011	0.079		
Unknown type of residence permit	-0.145	0.050	***	-0.288	0.037	***	
Age 16-24	1.221	0.061	***	1.132	0.048	***	
Age 25-34	1.068	0.050	***	0.799	0.037	***	
Age 35-44	0.803	0.048	***	0.475	0.034	***	
Children 3-6 years of age	0.299	0.038	***	0.109	0.032	***	
Children 7-17 years of age	0.544	0.042	***	0.214	0.035	***	
No children	0.415	0.039	***	0.035	0.034		
Single	-0.163	0.029	***	-0.095	0.031	***	
Working experience up to 1 year	1.675	0.035	***	1.296	0.028	***	
Working experience 1-3 years	1.860	0.044	***	1.667	0.033	***	
Working experience 3 years or more	1.776	0.058	***	1.779	0.042	***	
Lives in provincial town municipality	-0.078	0.032	***	0.026	0.024		
Lives in a rural district	0.111	0.045	***	0.167	0.037	***	
Danish education	-1.536	0.142	***	-1.280	0.114	***	
Years of Danish education	0.178	0.012	***	0.126	0.010	***	
Foreign education	-0.420	0.087	***	-0.387	0.078	***	
Years of foreign education	0.052	0.007	***	0.044	0.006	***	
No visits to doctors	-0.163	0.043	***	0.024	0.025		
5-9 visits to doctors	-0.083	0.031	***	-0.152	0.025	***	
10-19 visits to doctors	-0.346	0.035	***	-0.512	0.033	***	
20 or more visits to doctors	-0.651	0.054	***	-0.899	0.067	***	
Local unemployment rate x 10	-0.620	0.135	***	-0.689	0.107	***	
Social assistance spell began 1998	-0.075	0.044	*	-0.117	0.034	***	
Social assistance spell began 1999	-0.162	0.052	***	-0.211	0.040	***	
Social assistance spell began 2000	-0.047	0.058		-0.141	0.045	***	
Social assistance spell began 2001	-0.236	0.062	***	-0.207	0.049	***	
Social assistance spell began 2002	-0.267	0.064	***	-0.441	0.051	***	
Social assistance spell began 2003	-0.355	0.061	***	-0.415	0.048	***	
Duration 0-3 months	-5.432	0.153	***	-4.171	0.118	***	
Duration 3-6 months	-6.168	0.154	***	-4.931	0.120	***	
Duration 6-12 months	-6.406	0.152	***	-5.113	0.119	***	
Duration 12-24 months	-6.751	0.151	***	-5.410	0.118	***	
Duration 24- months	-7.033	0.150	***	-5.811	0.118	***	

Table A2. Estimated parameters for controls in the hazard rate to employment

Note. \*\*\*, \*\* and \* indicate significance at the 1%, 5% and 10% level, respectively. The last five variables are the duration dependent constant terms. See the note to Table A 1 for reference categories of the explanatory variables.

		Females			Males		
Quartile	First	Second	Third	First	Second	Third	
	Duration until start of programme (in months)						
Employment with wage subsidy	2	5	11	2	5	11	
Direct employment programme	2	7	19	2	5	13	
Other programmes	3	11	26	2	6	16	
		Duration of programme (in months)					
Employment with wage subsidy	3	5	7	3	5	7	
Direct employment programme	2	5	7	2	3	6	
Other programmes	2	4	9	2	3	6	

Table A3. Characteristics of ALMP spells used to calculate marginal effects: Quartiles in the distributions of duration of social assistance spells until start of programme and time spent in the programme



Figure 1. Kaplan-Meier survival functions (survival until regular employment)

#### Endnotes

- <sup>1</sup> The timing-of-events duration model has been used to evaluate ALMPs in several previous studies; see e.g., Richardson & van den Berg (2001), Bolvig *et al.* (2003), van den Berg *et al.* (2004), Abbring *et al.* (2005), Crépon *et al.* (2005), Lalive *et al.* 2005, 2008), Clausen *et al.* (2009) and Rosholm & Svarer (2008).
  - <sup>2</sup> Non-western countries are defined as countries which are not Nordic, not in the EU (as of May 2004), and not the US, Canada, Australia, New Zealand, Switzerland, Andorra, Liechtenstein, Monaco or San Marino.

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## The Effects of Active Labour Market Policies for Immigrants Receiving Social Assistance in Denmark

We estimate the effect of active labour market programmes on the exit rate to regular employment for non-western immigrants in Denmark who receive social assistance. We use the timing-of-events duration model and rich administrative data. We find large positive post-programme effects, and, surprisingly, even most in-programme effects are positive. The effects are largest for subsidized employment programmes, but effects are also large and significant for direct employment programmes and other programmes. The effects are larger if programmes begin after six months of unemployment. Implications of our estimates are illustrated by calculating effects on the duration to regular employment over a five-year period.

